

Principal Component Analysis

Linear Least Squares Approximation

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Definition (point set case)

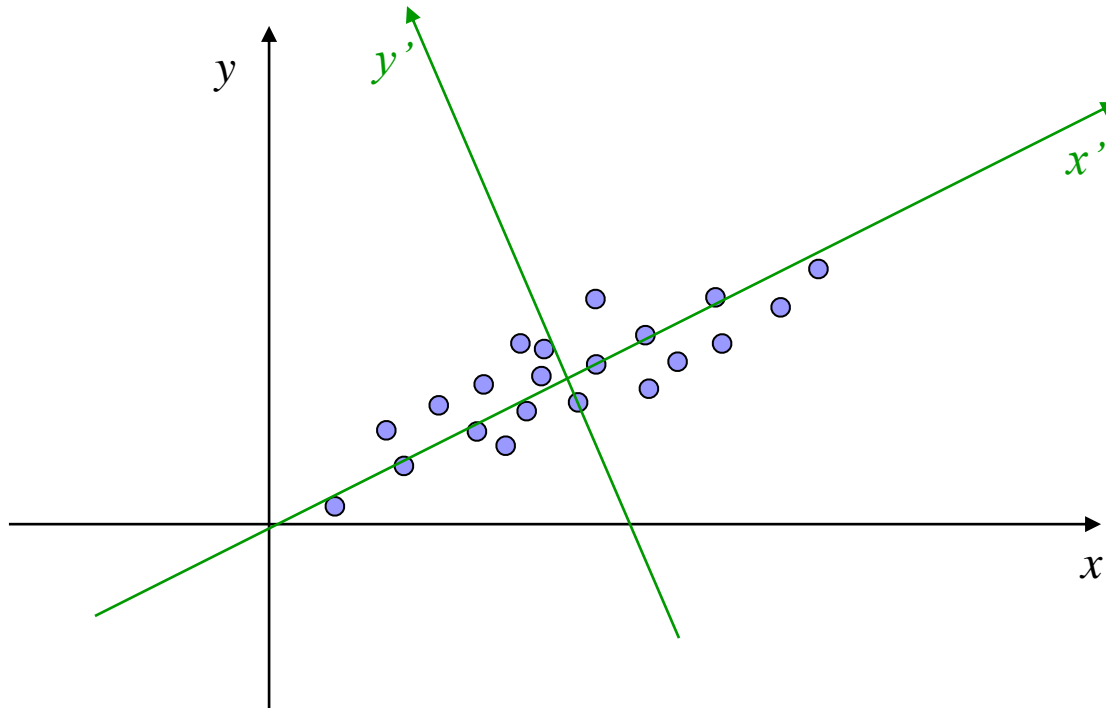
- Given a point set $\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_n \in R^d$, linear least squares fitting amounts to find the linear sub-space of R^d which **minimizes the sum of squared distances** from the points to their projection onto this linear sub-space.

Definition (point set case)

- This problem is equivalent to search for the linear sub-space which **maximizes the variance of projected points**, the latter being obtained by eigen decomposition of the covariance (scatter) matrix.
- Eigenvectors corresponding to large eigenvalues are the directions in which the data has **strong component**, or equivalently large variance. If eigenvalues are the same there is no preferable sub-space.

PCA – the general idea

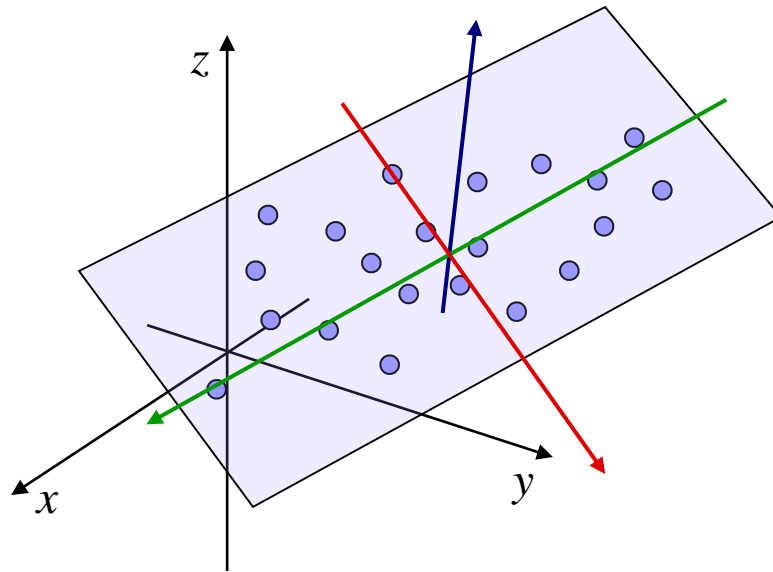
- PCA finds an orthogonal basis that best represents given data set.



- The sum of distances² from the x' axis is minimized.

PCA – the general idea

- PCA finds an orthogonal basis that best represents given data set.



3D point set

- PCA finds a best approximating plane (again, in terms of $\sum distances^2$)

Notations

- Denote our data points by $\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_n \in R^d$

$$\mathbf{x}_1 = \begin{pmatrix} x_1^1 \\ x_1^2 \\ \vdots \\ x_1^d \end{pmatrix}, \mathbf{x}_2 = \begin{pmatrix} x_2^1 \\ x_2^2 \\ \vdots \\ x_2^d \end{pmatrix}, \dots, \mathbf{x}_n = \begin{pmatrix} x_n^1 \\ x_n^2 \\ \vdots \\ x_n^d \end{pmatrix}$$

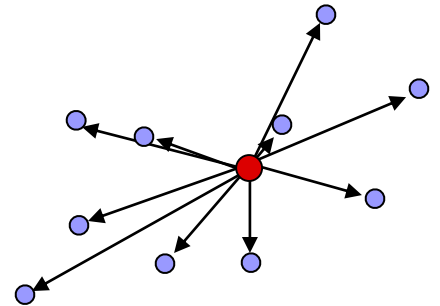
The origin of the new axes

- The origin is zero-order approximation of our data set (a point)
- It will be the center of mass:

$$\mathbf{m} = \frac{1}{n} \sum_{i=1}^n \mathbf{x}_i$$

- It can be shown that:

$$\mathbf{m} = \operatorname{argmin}_{\mathbf{x}} \sum_{i=1}^n \|\mathbf{x}_i - \mathbf{x}\|^2$$



Scatter matrix

- Denote $\mathbf{y}_i = \mathbf{x}_i - \mathbf{m}$, $i = 1, 2, \dots, n$

$$S = YY^T$$

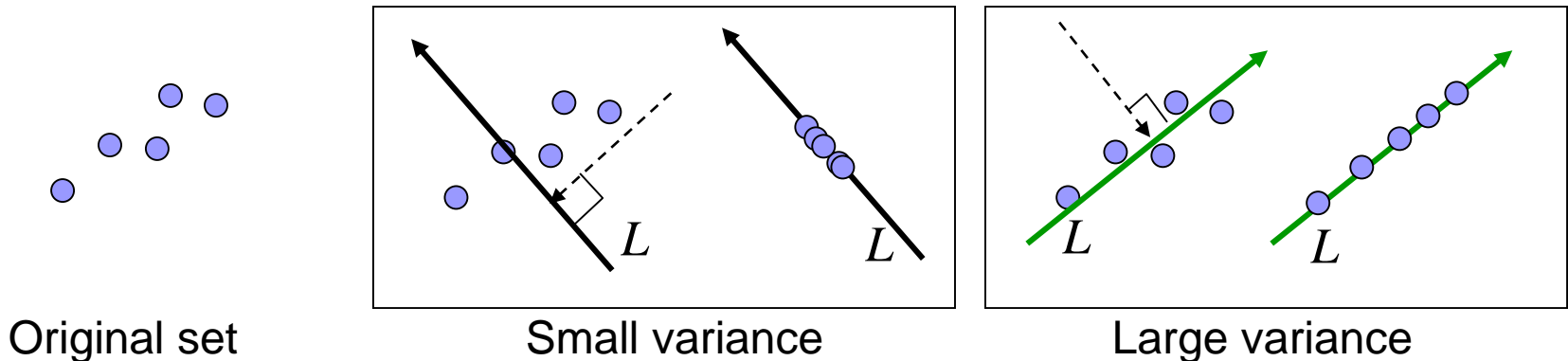
where Y is $d \times n$ matrix with \mathbf{y}_k as columns ($k = 1, 2, \dots, n$)

$$S = \underbrace{\begin{pmatrix} \mathbf{y}_1^1 & \mathbf{y}_2^1 & \cdots & \mathbf{y}_n^1 \\ \mathbf{y}_1^2 & \mathbf{y}_2^2 & \cdots & \mathbf{y}_n^2 \\ \vdots & \vdots & \cdots & \vdots \\ \mathbf{y}_1^d & \mathbf{y}_2^d & \cdots & \mathbf{y}_n^d \end{pmatrix}}_Y \underbrace{\begin{pmatrix} \mathbf{y}_1^1 & \mathbf{y}_1^2 & \cdots & \mathbf{y}_1^d \\ \mathbf{y}_2^1 & \mathbf{y}_2^2 & \cdots & \mathbf{y}_2^d \\ \vdots & \vdots & \cdots & \vdots \\ \mathbf{y}_n^1 & \mathbf{y}_n^2 & \cdots & \mathbf{y}_n^d \end{pmatrix}}_{Y^T}$$

Variance of projected points

- In a way, S measures variance (= scatterness) of the data in different directions.
- Let's look at a line L through the center of mass \mathbf{m} , and project our points \mathbf{x}_i onto it. The variance of the projected points \mathbf{x}'_i is:

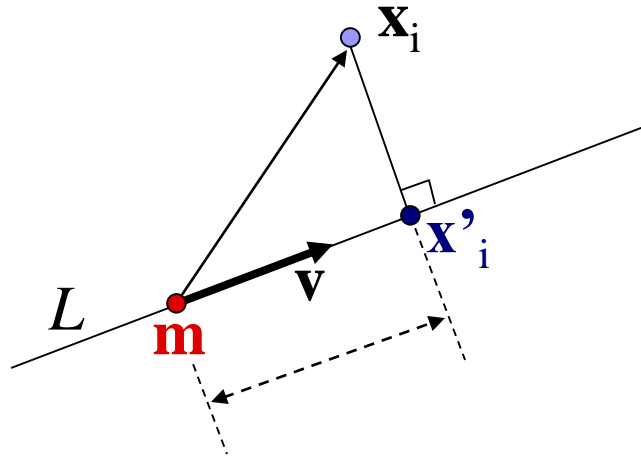
$$\text{var}(L) = \frac{1}{n} \sum_{i=1}^n \|\mathbf{x}'_i - \mathbf{m}\|^2$$



Variance of projected points

- Given a direction \mathbf{v} , $\|\mathbf{v}\| = 1$, the projection of \mathbf{x}_i onto $L = \mathbf{m} + \mathbf{v}t$ is:

$$\|\mathbf{x}'_i - \mathbf{m}\| = \langle \mathbf{v}, \mathbf{x}_i - \mathbf{m} \rangle \|\mathbf{v}\| = \langle \mathbf{v}, \mathbf{y}_i \rangle = \mathbf{v}^T \mathbf{y}_i$$



Variance of projected points

- So,

$$\begin{aligned}\text{var}(L) &= \frac{1}{n} \sum_{i=1}^n \|\mathbf{x}'_i - \mathbf{m}\|^2 = \frac{1}{n} \sum_{i=1}^n (\mathbf{v}^T \mathbf{y}_i)^2 = \frac{1}{n} \|\mathbf{v}^T \mathbf{Y}\|^2 = \\ &= \frac{1}{n} \|\mathbf{Y}^T \mathbf{v}\|^2 = \frac{1}{n} \langle \mathbf{Y}^T \mathbf{v}, \mathbf{Y}^T \mathbf{v} \rangle = \frac{1}{n} \mathbf{v}^T \mathbf{Y} \mathbf{Y}^T \mathbf{v} = \frac{1}{n} \mathbf{v}^T \mathbf{S} \mathbf{v} = \frac{1}{n} \langle \mathbf{S} \mathbf{v}, \mathbf{v} \rangle\end{aligned}$$

$$\sum_{i=1}^n (\mathbf{v}^T \mathbf{y}_i)^2 = \sum_{i=1}^n \left((v^1 \ v^2 \ \dots \ v^d) \begin{pmatrix} y_i^1 \\ y_i^2 \\ \vdots \\ y_i^d \end{pmatrix} \right)^2 = \left\| (v^1 \ v^2 \ \dots \ v^d) \begin{pmatrix} y_1^1 & y_2^1 & \dots & y_n^1 \\ y_1^2 & y_2^2 & & y_n^2 \\ \vdots & \vdots & & \vdots \\ y_1^d & y_2^d & \dots & y_n^d \end{pmatrix} \right\|^2 = \|\mathbf{v}^T \mathbf{Y}\|^2$$

Directions of maximal variance

- So, we have: $\text{var}(L) = \langle S\mathbf{v}, \mathbf{v} \rangle$

- Theorem:

Let $f: \{\mathbf{v} \in R^d \mid \|\mathbf{v}\| = 1\} \rightarrow R$,

$$f(\mathbf{v}) = \langle S\mathbf{v}, \mathbf{v} \rangle \quad (\text{and } S \text{ is a symmetric matrix}).$$

Then, the extrema of f are attained at the eigenvectors of S .

- So, eigenvectors of S are directions of maximal/minimal variance.

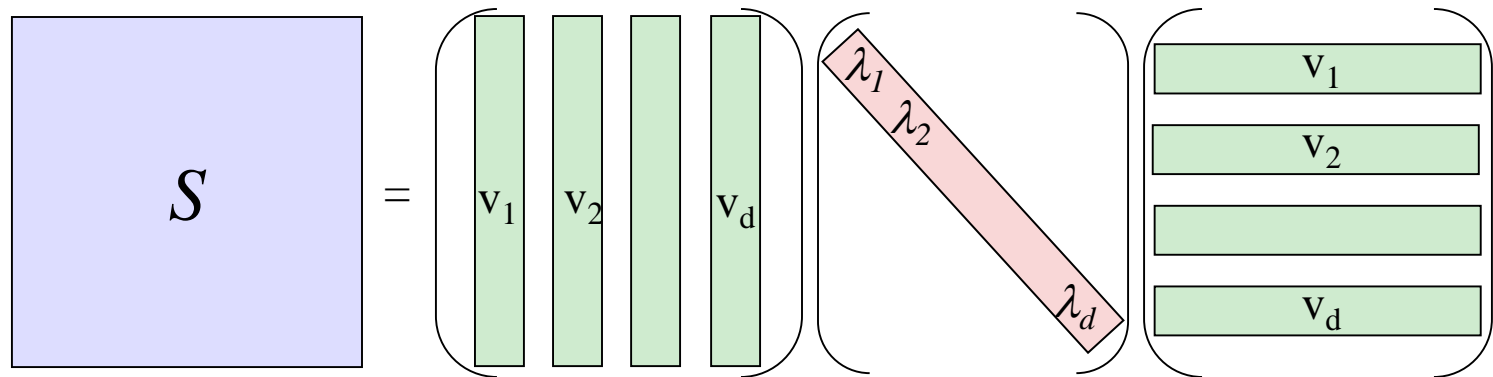
Summary so far

- We take the centered data points $\mathbf{y}_1, \mathbf{y}_2, \dots, \mathbf{y}_n \in \mathbb{R}^d$
- Construct the scatter matrix $S = YY^T$
- S measures the variance of the data points
- Eigenvectors of S are directions of max/min variance.

Scatter matrix - eigendecomposition

- S is symmetric

⇒ S has eigendecomposition: $S = V\Lambda V^T$

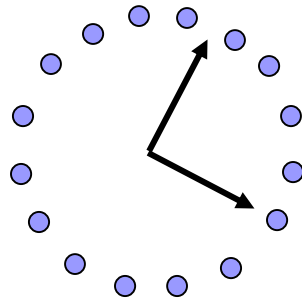


The eigenvectors form
orthogonal basis

Principal components

- Eigenvectors that correspond to **big** eigenvalues are the directions in which the data has strong components (= large variance).
- If the eigenvalues are more or less the same – there is no preferable direction.

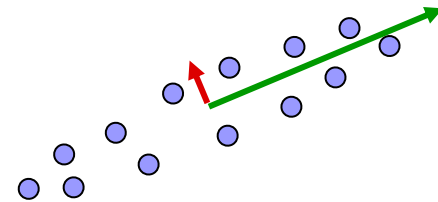
Principal components



- There's no preferable direction
- S looks like this:

$$\begin{pmatrix} \lambda & \\ & \lambda \end{pmatrix}$$

- Any vector is an eigenvector

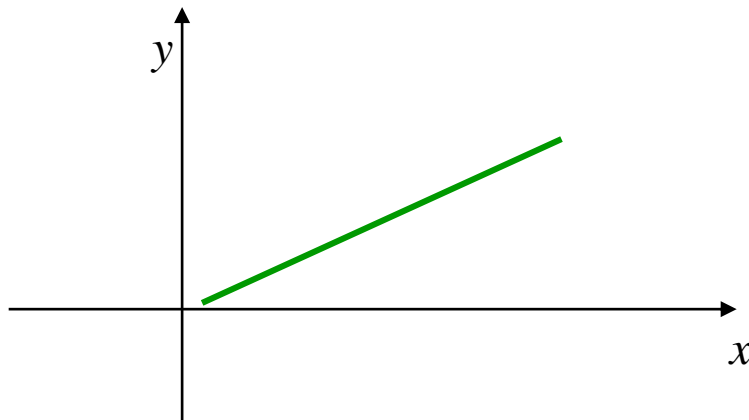
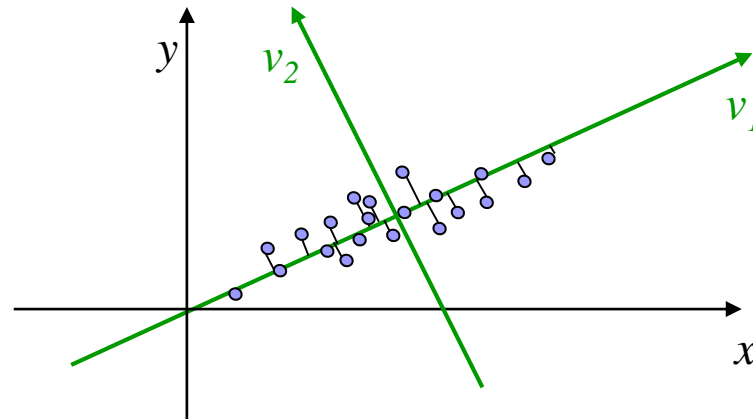


- There is a clear preferable direction
- S looks like this:

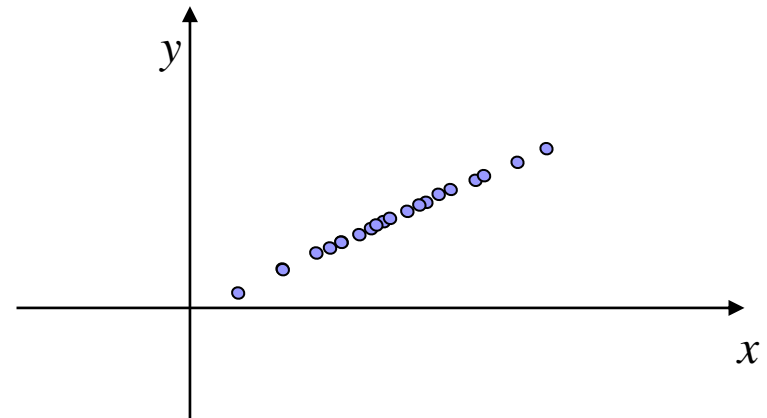
$$V \begin{pmatrix} \lambda & \\ & \mu \end{pmatrix} V^T$$

- μ is close to zero, much smaller than λ .

For approximation



This line segment approximates the original data set



The projected data set approximates the original data set

For approximation

- In general dimension d , the eigenvalues are sorted in descending order:

$$\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_d$$

- The eigenvectors are sorted accordingly.
- To get an approximation of dimension $d' < d$, we take the d' first eigenvectors and look at the subspace they span ($d' = 1$ is a line, $d' = 2$ is a plane...)

For approximation

- To get an approximating set, we project the original data points onto the chosen subspace:

$$\mathbf{x}_i = \mathbf{m} + \alpha_1 \mathbf{v}_1 + \alpha_2 \mathbf{v}_2 + \dots + \alpha_{d'} \mathbf{v}_{d'} + \dots + \alpha_d \mathbf{v}_d$$

Projection:

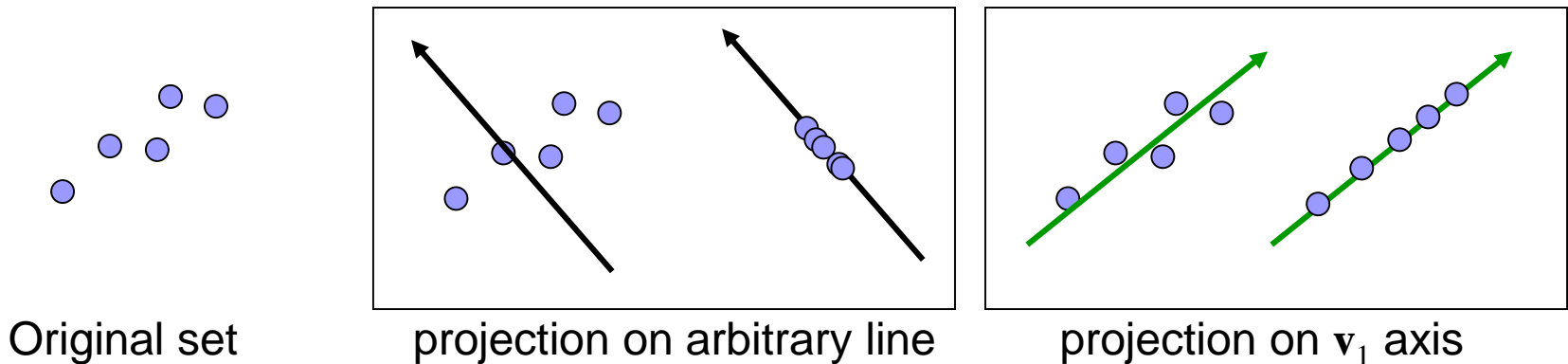
$$\mathbf{x}_i' = \mathbf{m} + \underbrace{\alpha_1 \mathbf{v}_1 + \alpha_2 \mathbf{v}_2 + \dots + \alpha_{d'} \mathbf{v}_{d'}}_{\text{subspace}} + 0 \cdot \mathbf{v}_{d'+1} + \dots + 0 \cdot \mathbf{v}_d$$

Optimality of approximation

- The approximation is optimal in **least-squares sense**. It gives the minimal of:

$$\sum_{k=1}^n \|\mathbf{x}_k - \mathbf{x}'_k\|^2$$

- The projected points have maximal variance.

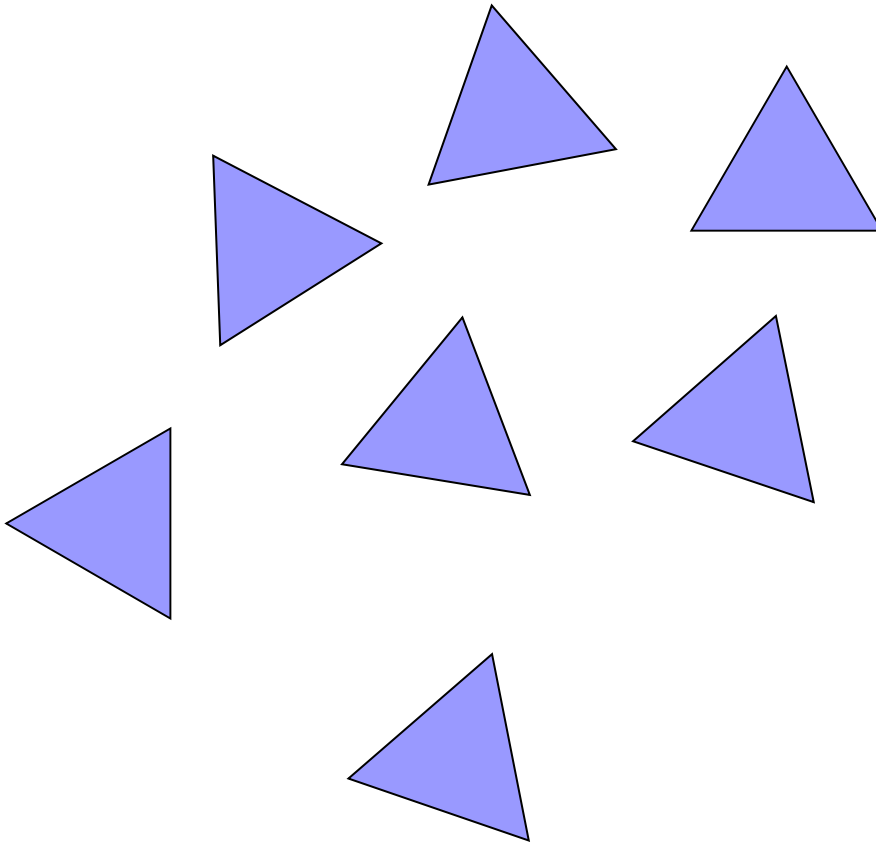


PCA on Point Sets

$$S = \begin{pmatrix} y_1^1 & y_2^1 & \cdots & y_n^1 \\ y_1^2 & y_2^2 & & y_n^2 \\ \vdots & \vdots & & \vdots \\ y_1^d & y_2^d & \cdots & y_n^d \end{pmatrix} \begin{pmatrix} y_1^1 & y_1^2 & \cdots & y_1^d \\ y_2^1 & y_2^2 & \cdots & y_2^d \\ \vdots & & & \vdots \\ y_n^1 & y_n^2 & \cdots & y_n^d \end{pmatrix}$$

[demo](#)

PCA on Geometric Primitives?



$$C_i = \int_{s_i} x x^T dx$$

Coordinate relative to center of mass